

Let s^2 be the sample variance of a random sample of size 6 from the normal distribution $N(\mu, 12)$. Find $P(2.76 < s^2 < 30.80)$

$$\frac{(n-1)s^2}{\sigma^2} \sim \chi^2(n-1) \quad \frac{ss^2}{12} \sim \chi^2(5)$$

$$P(2.30 \leq s^2 \leq 22.5) \leq P\left(\frac{5}{12}(2.76) \leq \chi^2(5) \leq \frac{5}{12}(30.80)\right)$$

$$= P(1.15 \leq \chi^2(5) \leq 12.8331) = 0.95 - 0.025 = 0.925$$

$$P(\chi^2(5) \geq 12.8325) = 0.025$$

$$P(\chi^2(5) \geq 1.15) = 0.95$$

Let $X_1, \dots, X_5$ be a random sample of size $n=5$ from $N(0, \sigma^2)$

(a) Find the constant c so that $\frac{c(X_1 - X_2)}{\sqrt{X_3^2 + X_4^2 + X_5^2}}$ has a t -distribution

(b) How many degrees of freedom are associated with the T ?

$$\bar{X}_1 - \bar{X}_2 \sim N(0, 2\sigma^2) \quad \frac{X_3^2 + X_4^2 + X_5^2}{\sigma^2} \sim \chi^2(3)$$

$$\frac{\frac{X_1 - X_2}{2\sigma^2}}{\sqrt{\frac{X_3^2 + X_4^2 + X_5^2}{3\sigma^2}}} = \frac{\sqrt{3}}{2} \frac{X_1 - X_2}{\sqrt{X_3^2 + X_4^2 + X_5^2}} \sim t(3)$$

Extra

Let $\bar{Y}$ be the sample mean of a random sample of size S from a normal distribution with mean $\mu=0$ and $\sigma^2=125$. Determine c so that $P(\bar{Y} \leq c) = 0.90$

$$E[\bar{Y}] = \mu \quad \text{Var}(\bar{Y}) = \frac{\sigma^2}{n} = \frac{125}{5} = 25$$

$$0.10 = P(N(0,1) \geq 2.326)$$

$$P(\bar{Y} \geq c) = P\left(\frac{\bar{Y} - \mu}{\sqrt{\sigma^2/n}} \geq \frac{c}{5}\right)$$

$$2.326 = \frac{c}{5} \quad c = 5(2.326) = 11.63$$

Let $X_1, \dots, X_{25}$ and let $Y_1, \dots, Y_{25}$ be two independent random samples from two normal distributions $N(0, 16)$ and $N(1, 9)$, respectively

Let $\bar{X}$ and let $\bar{Y}$ denote the corresponding sample means.

Compute $P(\bar{X} > \bar{Y})$

$$E[\bar{X}] = 0 \quad \text{Var}(\bar{X}) = \frac{16}{25}$$

$$E[\bar{Y}] = 1 \quad \text{Var}(\bar{Y}) = \frac{9}{25}$$

$$E[\bar{X} - \bar{Y}] = -1 \quad \text{Var}(\bar{X} - \bar{Y}) = \frac{16}{25} + \frac{9}{25} = 1$$

$$P(\bar{X} > \bar{Y}) = P(\bar{X} - \bar{Y} > 0) = P(N(0, 1) > 1) = 0.1587$$

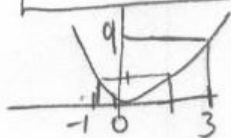
Let X_1, X_2 be a random sample from a distribution having the pdf
 $f(x) = e^{-x}, 0 < x < \infty$, zero elsewhere. Show that $Z = \frac{X_1}{X_2}$ has
 an F distribution

$$Y \sim \chi^2(k) \quad f(y) = \frac{y^{\frac{k}{2}-1} e^{-\frac{y}{2}}}{\Gamma(\frac{k}{2}) 2^{k/2}}$$

$$2X_1 \sim \chi^2(2), \quad 2X_2 \sim \chi^2(2)$$

$$\text{So } Z = \frac{\frac{2X_1}{2}}{\frac{2X_2}{2}} = \frac{X_1}{X_2} \sim F(2, 2)$$

If X has the pdf $f(x) = \frac{1}{4}, -1 < x < 3, f(x) = 0$ otherwise
 Find the pdf of $Y = X^2$



$$\text{For } 0 < y < 1, \quad P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} < X < \sqrt{y}) = \frac{2\sqrt{y}}{4} = \frac{\sqrt{y}}{2}$$

$$\text{For } 1 < y < 9, \quad P(Y < y) = P(-1 < X < \sqrt{y}) = \frac{\sqrt{y} + 1}{4}$$

$$F_Y(y) = \begin{cases} 0 & y < 0 \\ \frac{\sqrt{y}}{2} & 0 \leq y < 1 \\ \frac{\sqrt{y} + 1}{4} & 1 \leq y < 9 \\ 0 & y \geq 9 \end{cases}$$

$$f_Y(y) = \begin{cases} \frac{1}{4\sqrt{y}} & 0 < y \leq 1 \\ \frac{1}{8\sqrt{y}} & 1 < y < 9 \\ 0 & \text{else} \end{cases}$$

Let X_1, X_2, X_3 be three independent chi-square variables with ν_1, ν_2, ν_3 degrees of freedom respectively

(a) Show that $Y_1 = \frac{X_1}{X_2}$ and $Y_2 = X_1 + X_2$ are independent.

(b) Deduce that

$$\frac{X_1/\nu_1}{X_2/\nu_2} \quad \text{and} \quad \frac{X_3/\nu_3}{(X_1+X_2)/(\nu_1+\nu_2)}$$

are independent Fr-variables

$$(y_1, y_2) = h(x_1, x_2), \quad (x_1, x_2) = h^{-1}(y_1, y_2)$$

The joint pdf of Y_1 and Y_2 is

$$f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}(h^{-1}(y_1, y_2)) |J h^{-1}(y_1, y_2)|$$

$$X_1 = X_2 Y_1 = (Y_2 - X_1) Y_1 = Y_1 Y_2 - X_1 Y_1 \quad X_1 = \frac{Y_1 Y_2}{1+Y_1} = Y_2 - \frac{Y_2}{1+Y_1}$$

$$X_2 = \frac{X_1}{Y_1} = \frac{Y_2}{1+Y_1} \quad J = \det \begin{pmatrix} \frac{Y_2}{(1+Y_1)^2} & \frac{Y_1}{1+Y_1} \\ -\frac{Y_2}{(1+Y_1)^2} & \frac{1}{1+Y_1} \end{pmatrix}$$

$$= \frac{Y_2}{(1+Y_1)^3} + \frac{Y_1 Y_2}{(1+Y_1)^3} = \frac{Y_2}{(1+Y_1)^2}$$

$$f_{X_1, X_2}(x_1, x_2) = \frac{x_1^{\frac{\nu_1}{2}-1} e^{-\frac{x_1}{2}}}{\Gamma(\frac{\nu_1}{2}) 2^{\nu_1/2}} \frac{x_2^{\frac{\nu_2}{2}-1} e^{-\frac{x_2}{2}}}{\Gamma(\frac{\nu_2}{2}) 2^{\nu_2/2}}, \quad x_1, x_2 > 0$$

$$f_{Y_1, Y_2}(y_1, y_2) = \frac{\left(\frac{y_1 y_2}{1+y_1}\right)^{\frac{v_1}{2}-1} e^{-\frac{y_1 y_2}{2(1+y_1)}}}{\Gamma(\frac{v_1}{2}) 2^{v_1/2}} \frac{\left(\frac{y_2}{1+y_1}\right)^{\frac{v_2}{2}-1} e^{-\frac{y_2}{2(1+y_1)}}}{\Gamma(\frac{v_2}{2}) 2^{v_2/2}} \frac{y_2}{(1+y_1)^2}$$

$$= \frac{y_1^{\frac{v_1}{2}-1}}{\Gamma(\frac{v_1}{2}) 2^{v_1/2} \Gamma(\frac{v_2}{2}) 2^{v_2/2} (1+y_1)^{\frac{v_1+v_2}{2}}} \frac{y_2^{\frac{v_2}{2}-1} e^{-\frac{y_2}{2(1+y_1)}}}{y_2^{\frac{v_1+v_2}{2}-1} e^{-\frac{y_2}{2(1+y_1)}}}, y_1, y_2 \geq 0$$

$$Y_1 \sim F(v_1, v_2) \quad Y_2 \sim X^2(v_1+v_2)$$

Let X be a r.v. with pdf $f(x) = 2x e^{-x^2}$, $x \geq 0$, elsewhere 0.
Find the pdf of $Y = X^2$ $x = h^{-1}(y) = \sqrt{y}$

$$f_Y(y) = f_X(h^{-1}(y)) \left| \frac{dh^{-1}(y)}{dy} \right| = 2\sqrt{y} e^{-y} \frac{1}{2\sqrt{y}} = e^{-y}, y \geq 0,$$

Let X_1, X_2 be independent r.v.'s with pdf $f(x) = e^{-x}$, $x \geq 0$, elsewhere 0.

Find the pdf of $Y = X_1 + X_2$

$$X_1 \sim \text{Gamma}(1, 1) \quad X_2 \sim \text{Gamma}(1, 1)$$

$$Y = X_1 + X_2 \sim \text{Gamma}(2, 1)$$

$$f_Y(y) = y e^{-y}, y \geq 0$$

Let X_1 and X_2 have the joint pdf $f_{X_1, X_2}(x_1, x_2) = 2e^{-x_1 - x_2}$, $0 < x_1 < x_2 < \infty$, zero elsewhere. Find the joint pdf of $Y_1 = 2X_1$ and $Y_2 = X_2 - X_1$ and argue that Y_i are independent

$$\begin{aligned} Y_1 = 2X_1, \quad Y_2 = X_2 - X_1 & \begin{cases} X_1 = Y_1/2 \\ X_2 = \frac{Y_1}{2} + Y_2 \end{cases} \end{aligned}$$

$$h'(Y_1, Y_2) = \left(\frac{Y_1}{2}, \frac{Y_1}{2} + Y_2 \right)$$

$$J h' = \det \begin{pmatrix} \frac{1}{2} & 0 \\ \frac{1}{2} & 1 \end{pmatrix} = \frac{1}{2}$$

$$f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}\left(\frac{y_1}{2}, \frac{y_1}{2} + y_2\right) \frac{1}{2} = e^{-y_1 + y_2}$$

$$0 < y_1, \quad 0 < y_2$$

$$0 < \frac{y_1}{2} < \frac{y_1}{2} + y_2$$

$$f_{Y_1}(y_1) = e^{-y_1}, \quad y_1 > 0, \quad f_{Y_2}(y_2) = e^{-y_2}, \quad y_2 \geq 0$$

Y_1, Y_2 are independent