

Definition

Let $Z \sim N(0, 1)$ and let $W \sim \chi^2(\nu)$

If Z and W are independent, then

$$T = \frac{Z}{\sqrt{\frac{W}{\nu}}} \sim t(\nu)$$

$$f(t) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\pi\nu} \Gamma\left(\frac{\nu}{2}\right)} \frac{1}{\left(1 + \frac{t^2}{\nu}\right)^{\frac{\nu+1}{2}}}, \quad -\infty < t < \infty$$

$$E[T] = 0 \quad V(T) = \frac{\nu}{\nu-2}, \quad \text{if } \nu > 2.$$

$$\left. \begin{array}{l} \frac{\sqrt{n}}{\sigma} (\bar{Y} - \mu) \sim N(0, 1) \\ \frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1) \end{array} \right\} t = \frac{\frac{\sqrt{n}(\bar{Y} - \mu)}{\sigma}}{\sqrt{\frac{(n-1)S^2}{\sigma^2} \cdot \frac{1}{n-1}}} = \frac{\sqrt{n}(\bar{Y} - \mu)}{S} \sim t(n-1)$$

If t_{α} is the value such that $P(T \geq t_{\alpha}) = \frac{\alpha}{2}$

$$P\left(-t_{\frac{\alpha}{2}} \leq \frac{\sqrt{n}(\bar{Y} - \mu)}{S} \leq t_{\frac{\alpha}{2}}\right) = 1 - \alpha$$

Proof Let $X \sim N(0, 1)$, $Y \sim \chi_p^2$, X and Y independent

$$f_X(x) = \frac{e^{-x^2/2}}{\sqrt{2\pi}} \quad f_Y(y) = \frac{y^{\frac{p}{2}-1} e^{-y/2}}{\Gamma(\frac{p}{2}) 2^{p/2}}, \quad y > 0$$

Let $U = \frac{X}{\sqrt{Y/p}}$, $V = Y$ $u \in \mathbb{R}, v > 0$
 $x \in \mathbb{R}, y > 0$

$$X = \frac{u\sqrt{v}}{\sqrt{p}} \quad Y = v$$

$$f_{u,v}(u,v) = f_{X,Y}(x(u,v), y(u,v)) \left| \frac{\partial(x,y)}{\partial(u,v)} \right|$$

$$\left| \frac{\partial(x,y)}{\partial(u,v)} \right| = \det \begin{pmatrix} \frac{\sqrt{v}}{\sqrt{p}} & \frac{u}{2\sqrt{v}p} \\ 0 & 1 \end{pmatrix} = \frac{\sqrt{v}}{\sqrt{p}}$$

$$f_{u,v}(u,v) = \frac{e^{-\frac{u^2 v}{2p}}}{\sqrt{2\pi}} \frac{v^{\frac{p}{2}-1} e^{-v/2}}{\Gamma(\frac{p}{2}) 2^{p/2}} \frac{\sqrt{v}}{\sqrt{p}} = \frac{e^{-\frac{u^2 v}{2p}} v^{\frac{p}{2}-1} e^{-v/2}}{\sqrt{2\pi p} \Gamma(\frac{p}{2}) 2^{p/2}}$$

$$f_u(u) = \int_0^\infty \frac{e^{-\frac{u^2 v}{2p}} v^{\frac{p}{2}-1} e^{-v/2}}{\sqrt{2\pi p} \Gamma(\frac{p}{2}) 2^{p/2}} dv = \int_0^\infty \frac{v^{\frac{p}{2}-1} e^{-v(\frac{1}{2} + \frac{u^2}{2p})}}{\sqrt{2\pi p} \Gamma(\frac{p}{2}) 2^{p/2}} dv$$

$$= \frac{\Gamma(\frac{p+1}{2}) \left(\frac{1}{\frac{1}{2} + \frac{u^2}{2p}} \right)^{\frac{p+1}{2}}}{\sqrt{2\pi p} \Gamma(\frac{p}{2}) 2^{p/2}} = \frac{\Gamma(\frac{p+1}{2}) \left(\frac{2p}{p+u^2} \right)^{\frac{p+1}{2}}}{\sqrt{2\pi p} \Gamma(\frac{p}{2}) 2^{p/2}}$$

$$= \frac{\Gamma(\frac{p+1}{2})}{\sqrt{\pi p} \Gamma(\frac{p}{2})} \frac{1}{\left(1 + \frac{u^2}{p}\right)^{\frac{p+1}{2}}} \quad - u > 0$$

7.12 Refer to Exercise 7.1. Suppose that in the forest fertilization problem of population standard deviation of basal areas is not known and must be estimated from the sample. If a random sample of $n=9$ basal areas is to be measured, find two statistics g_1 and g_2 such that $P(g_1 \leq \bar{Y} - \mu \leq g_2) = 0.90$

$$\frac{\sqrt{n}(\bar{Y} - \mu)}{S} \sim t(8) \quad P(-1.86 \leq t(8) \leq 1.86) = 0.90$$

$t_{0.050}$

$\downarrow$
1.86 $\leftarrow$ 8 d.f.

$$P(-1.86 \leq \frac{\sqrt{9}(\bar{Y} - \mu)}{S} \leq 1.86)$$

$$P\left(-\frac{1.86}{3} \leq \bar{Y} - \mu \leq \frac{1.86 \cdot 3}{3}\right)$$

7.14 Show

$$(i) E[T] = 0 \quad \forall \nu > 1$$

$$(ii) \text{Var}(T) = \frac{\nu}{\nu-2} \quad \forall \nu > 2$$

$$(a) E[T] = \sqrt{\nu} E[Z] E\left[\frac{1}{\sqrt{W}}\right]$$

$$E[Z] = 0$$

$$E\left[\frac{1}{\sqrt{W}}\right] = \int_0^{\infty} \frac{1}{w^{1/2}} \frac{w^{\frac{\nu}{2}-1} e^{-w/2}}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} dw$$

$$\frac{\nu}{2} - 1 - \frac{1}{2} > -1 \quad \text{or} \quad \nu > 1$$

$$(b) E[T^2] = \nu E[Z^2] E\left[\frac{1}{W}\right] = \nu \cdot 1 \cdot \frac{1}{\nu-2} = \frac{\nu}{\nu-2}$$

$$E\left[\frac{1}{W}\right] = \int_0^{\infty} \frac{1}{w} \frac{w^{\frac{\nu}{2}-1} e^{-w/2}}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} dw = \int_0^{\infty} \frac{w^{\frac{\nu}{2}-2} e^{-w/2}}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} dw$$

$$= \frac{\Gamma(\frac{\nu-2}{2}) 2^{\frac{\nu-2}{2}}}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} = \left(\frac{\nu-2}{2}\right) \cdot 2 = \nu-2$$