

## Large sample confidence intervals

Let  $\hat{\theta}$  be an estimator of  $\theta$ . Let  $\hat{\sigma}_{\hat{\theta}}^2(\theta) = \text{Var}_{\theta}(\hat{\theta})$  be the variance of  $\hat{\theta}$ . For many estimators,  $z = \frac{\hat{\theta} - \theta}{\hat{\sigma}_{\hat{\theta}}(\theta)}$  has an asymptotically normal distribution. Since  $\theta$  is unknown and estimated by  $\hat{\sigma}_{\hat{\theta}}^2 = \hat{\sigma}_{\hat{\theta}}^2(\hat{\theta})$ , this can be used to obtain large sample confidence intervals.

$$\lim_{n \rightarrow \infty} P\left(-\frac{z_{\alpha/2}}{\sqrt{\hat{\sigma}_{\hat{\theta}}^2}} \leq \frac{\hat{\theta} - \theta}{\sqrt{\hat{\sigma}_{\hat{\theta}}^2}} \leq \frac{z_{\alpha/2}}{\sqrt{\hat{\sigma}_{\hat{\theta}}^2}}\right) = 1-\alpha$$

An asymptotic  $1-\alpha$  confidence interval for  $\theta$  is  $\hat{\theta} \pm \frac{z_{\alpha/2}}{\sqrt{\hat{\sigma}_{\hat{\theta}}^2}}$

$$\hat{\theta} - \frac{z_{\alpha/2}}{\sqrt{\hat{\sigma}_{\hat{\theta}}^2}} \leq \theta \leq \hat{\theta} + \frac{z_{\alpha/2}}{\sqrt{\hat{\sigma}_{\hat{\theta}}^2}}$$

## Binomial confidence intervals

Let  $Y$  be a r.v. with a binomial distribution with parameters  $n$  and  $p$ ,  $n$  is known,  $p$  is unknown.

$$E[Y] = np \quad \text{Var}(Y) = np(1-p)$$

By the CLT,  $\frac{Y-np}{\sqrt{np(1-p)}}$  has an asymptotic normal distribution

$$\text{To estimate } p, \text{ we use } \hat{p} = \frac{y}{n}, \quad E[\hat{p}] = p, \quad \text{Var}(\hat{p}) = \frac{p(1-p)}{n}$$

$$\text{We estimate } \sigma_{\hat{p}} \text{ by } \hat{\sigma}_{\hat{p}} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

By the CLT  $\frac{\hat{p}-p}{\sqrt{\frac{\hat{p}(1-\hat{p})}{n}}}$  has an asymptotic normal distribution

$$\text{So, } \hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \text{ is an asymptotic } 1-\alpha \text{ confidence interval}$$

for  $p$ .

8.45 A New York Times/CBS survey of 224 children of ages 9 to 17 was conducted shortly after the January 28, 1986 space shuttle disaster. Two thirds of the children said that they would like to travel in space

(a) Construct a 90% confidence interval for the proportion of children ages 9 to 17 in 1986 who would like to experience space travel

(b) Do you think that "most" children in this age group think that they would like to experience space travel? Why?

(a)  $y = \text{no of children out of sample of 224, who would like to experience space travel}$

$y \sim \text{Bin}(n=224, p)$ ,  $p$  unknown.

$$\hat{p} = \frac{y}{n} = \frac{2}{3} \quad \hat{\sigma}_{\hat{p}} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = \sqrt{\frac{1}{224} \cdot \frac{2}{3} \cdot \frac{1}{3}} = \sqrt{\frac{1}{1008}}$$

A 90% confidence interval for  $p$  is

$$\hat{p} \pm 2 \cdot \hat{\sigma}_{\hat{p}} = \frac{2}{3} \pm (1.645) \frac{1}{\sqrt{1008}} = 0.6666 \pm 0.0518$$

$$= [0.6148, 0.7185]$$

(b) Yes, because the confidence interval is above 0.50

8.42 "Athletes at major colleges graduated, on the whole, at virtually the same rate as other students," according the NCAA. Suppose that, in a new poll of 500 athletes at major colleges the number graduating was 268.

Find a 98% confidence interval for  $p$ , the proportion of athletes at major colleges who graduate

$$\hat{p} = \frac{y}{n} = \frac{268}{500} = 0.536, \quad \sigma_{\hat{p}}^2 = \frac{\hat{p}(1-\hat{p})}{n} = \frac{1}{500} \left( \frac{268}{500} \left( 1 - \frac{268}{500} \right) \right)$$

$$= 0.0004974 \quad \sigma_{\hat{p}} = 0.0223$$

$$\hat{p} \pm 2 \frac{\sigma_{\hat{p}}}{\sqrt{n}} = 0.536 \pm 2.326(0.0223) = (0.4841, 0.5879)$$

$$\hat{p} \pm 2 \frac{\sigma_{\hat{p}}}{\sqrt{n}}$$

## Binomial confidence intervals - TWO samples

Assume that

$Y_1 \sim \text{Bin}(n_1, p_1)$ ,  $n_1$  known,  $p_1$  unknown

$Y_2 \sim \text{Bin}(n_2, p_2)$ ,  $n_2$  known,  $p_2$  unknown

$Y_1$  and  $Y_2$  are independent r.v.'s

We would like to estimate  $\theta = p_1 - p_2$

$$\text{Let } \hat{p}_1 = \frac{Y_1}{n_1}, \quad \hat{p}_2 = \frac{Y_2}{n_2}, \quad \hat{\theta} = \hat{p}_1 - \hat{p}_2$$

$\hat{\theta} = \hat{p}_1 - \hat{p}_2$  estimates  $p_1 - p_2$

$$E[\hat{p}_1 - \hat{p}_2] = p_1 - p_2$$

$$\text{Var}(\hat{p}_1 - \hat{p}_2) = \frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2} = \sigma_{\theta}^2$$

We estimate  $\sigma_{\theta}^2$  by

$$\hat{\sigma}_{\theta}^2 = \frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}$$

An asymptotic  $1-\alpha$  confidence interval for  $p_1 - p_2$  is

$$\hat{p}_1 - \hat{p}_2 - z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

8.51 For a comparison of the rates of defectives produced by two assembly lines, independent random samples of 100 items are selected from each line. Line A yielded 18 defectives in the sample, and line B yielded 12 defectives. Find a 98% confidence interval for the true difference in proportions of defectives for the two lines. Is there evidence here to suggest that one line produces a higher proportion of defectives than the other?

$Y_1$  = no. of defectives in Line A with probability of defective  $p_1$

$Y_2$  = no. of defectives in line B with probability of defective  $p_2$

$$\text{To estimate } p_1 - p_2 \text{ we use } \frac{Y_1}{n_1} - \frac{Y_2}{n_2} = \frac{18}{100} - \frac{12}{100} = 0.06$$

$$V(p_1 - p_2) = \frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2} = \frac{0.18(1-0.18)}{100} + \frac{0.12(1-0.12)}{100}$$

$$= 0.002532$$

$Z_{\frac{\alpha}{2}} = Z_{0.01} = 2.33$  The confidence interval is

$$0.06 \pm 2.33 \sqrt{0.002532} = 0.06 \pm 0.1172 = (-0.0572, 0.1772)$$

No, there is no evidence to suggest one line produces a higher proportion of defectives than the other